

Overview of Unmarked: An R Package for the Analysis of Data from Unmarked Animals

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Abstract

Unmarked aims to be a complete environment for the statistical analysis of data from surveys of unmarked animals. Currently, the focus is on hierarchical models that separately model a latent state (or states) and an observation process. This vignette provides a brief overview of the package — for a more thorough treatment see [2]

1 Overview of unmarked

Unmarked provides methods to estimate site occupancy, abundance, and density of animals (or possibly other organisms/objects) that cannot be detected with certainty. Numerous models are available that correspond to specialized survey methods such as temporally replicated surveys, distance sampling, removal sampling, and double observer sampling. These data are often associated with metadata related to the design of the study. For example, in distance sampling, the study design (line- or point-transect), distance class break points, transect lengths, and units of measurement need to be accounted for in the analysis. Unmarked uses S4 classes to store data and metadata in a way that allows for easy data manipulation, summarization, and model specification. Table 1 lists the currently implemented models and their associated fitting functions and data classes.

Model	Fitting Function	Data	Citation
Occupancy	occu	unmarkedFrameOccu	[4]
Royle-Nichols	occuRN	unmarkedFrameOccu	[8]
Point Count	pcount	unmarkedFramePCount	[6]
Distance-sampling	distsamp	unmarkedFrameDS	[7]
Generalized distance-sampling	gdistsamp	unmarkedFrameGDS	[1]
Arbitrary multinomial-Poisson	multinomPois	unmarkedFrameMPois	[5]
Colonization-extinction	colect	unmarkedMultFrame	[3]
Generalized multinomial-mixture	gmultmix	unmarkedFrameGMM	[5]

Table 1: Models handled by unmarked.

Each data class can be created with a call to the constructor function of the same name as described in the examples below.

2 Typical unmarked session

The first step is to import the data into R, which we do below using the **read.csv** function. Next, the data need to be formatted for use with a specific model fitting function. This can be accomplished with a call to the appropriate type of **unmarkedFrame**. For example, to prepare the data for a single-season site-occupancy analysis, the function **unmarkedFrameOccu** is used.

2.1 Importing and formatting data

```
> library(unmarked)
> wt <- read.csv(system.file("csv","widewt.csv", package="unmarked"))
```

```

> y <- wt[,2:4]
> siteCovs <- wt[,c("elev", "forest", "length")]
> obsCovs <- list(date=wt[,c("date.1", "date.2", "date.3")],
  ivel=wt[,c("ivel.1", "ivel.2", "ivel.3")])
> wt <- unmarkedFrameOccu(y = y, siteCovs = siteCovs, obsCovs = obsCovs)
> summary(wt)
unmarkedFrame Object

```

```

237 sites
Maximum number of observations per site: 3
Mean number of observations per site: 2.81
Sites with at least one detection: 79

```

Tabulation of y observations:

```

  0    1 <NA>
483 182   46

```

Site-level covariates:

elev	forest	length
Min. : -1.436125	Min. : -1.265352	Min. : 0.1823
1st Qu.: -0.940726	1st Qu.: -0.974355	1st Qu.: 1.4351
Median : -0.166666	Median : -0.064987	Median : 1.6094
Mean : 0.007612	Mean : 0.000088	Mean : 1.5924
3rd Qu.: 0.994425	3rd Qu.: 0.808005	3rd Qu.: 1.7750
Max. : 2.434177	Max. : 2.299367	Max. : 2.2407

Observation-level covariates:

date	ivel
Min. : -2.90434	Min. : -1.7533
1st Qu.: -1.11862	1st Qu.: -0.6660
Median : -0.11862	Median : -0.1395
Mean : -0.00022	Mean : 0.0000
3rd Qu.: 1.30995	3rd Qu.: 0.5493
Max. : 3.80995	Max. : 5.9795
NA's : 42	NA's : 46

Alternatively, the convenience function **csvToUMF** can be used

```

> wt <- csvToUMF(system.file("csv", "widewt.csv", package="unmarked"),
  long = FALSE, type = "unmarkedFrameOccu")

```

If not all sites have the same numbers of observations, then manual importation of data in long format can be tricky. **csvToUMF** seamlessly handles this situation.

```

> pcru <- csvToUMF(system.file("csv", "frog2001pcru.csv", package="unmarked"),
  long = TRUE, type = "unmarkedFrameOccu")

```

To help stabilize the numerical optimization algorithm, we recommend standardizing the covariates.

```

> obsCovs(pcru) <- scale(obsCovs(pcru))

```

2.2 Fitting models

Occupancy models can then be fit with the `occu()` function:

```

> fm1 <- occu(~1 ~1, pcru)
> fm2 <- occu(~ MinAfterSunset + Temperature ~ 1, pcru)
> fm2
Call:
occu(formula = ~MinAfterSunset + Temperature ~ 1, data = pcru)

```

Occupancy:

Estimate	SE	z	P(> z)
1.54	0.292	5.26	1.42e-07

Detection:

	Estimate	SE	z	P(> z)
(Intercept)	0.2098	0.206	1.017	3.09e-01
MinAfterSunset	-0.0855	0.160	-0.536	5.92e-01
Temperature	-1.8936	0.291	-6.508	7.60e-11

AIC: 356.7591

Here, we have specified that the detection process is modeled with the MinAfterSunset and Temperature covariates. No covariates are specified for occupancy here. See ?occu for more details.

2.3 Back-transforming parameter estimates

Unmarked fitting functions return unmarkedFit objects which can be queried to investigate the model fit. Variables can be back-transformed to the unconstrained scale using backTransform. Standard errors are computed using the delta method.

```
> backTransform(fm2, 'state')
Backtransformed linear combination(s) of Occupancy estimate(s)
```

Estimate	SE	LinComb	(Intercept)
0.823	0.0425	1.54	1

Transformation: logistic

The expected probability that a site was occupied is 0.823. This estimate applies to the hypothetical population of all possible sites, not the sites found in our sample. For a good discussion of population-level vs finite-sample inference, see Royle and Dorazio [9] page 117. Note also that finite-sample quantities can be computed in `unmarked` using empirical Bayes methods as demonstrated at the end of this document.

Back-transforming the estimate of ψ was easy because there were no covariates. Because the detection component was modeled with covariates, p is a function, not just a scalar quantity, and so we need to provide values of our covariates to obtain an estimate of p . Here, we request the probability of detection given a site is occupied and all covariates are set to 0.

```
> backTransform(linearComb(fm2, coefficients = c(1,0,0), type = 'det'))
Backtransformed linear combination(s) of Detection estimate(s)
```

Estimate	SE	LinComb	(Intercept)	MinAfterSunset	Temperature
0.552	0.051	0.21	1	0	0

Transformation: logistic

Thus, we can say that the expected probability of detection was 0.552 when time of day and temperature are fixed at their mean value. A predict method also exists, which can be used to obtain estimates of parameters at specific covariate values.

```
> newData <- data.frame(MinAfterSunset = 0, Temperature = -2:2)
> round(predict(fm2, type = 'det', newdata = newData, appendData=TRUE), 2)
Predicted SE lower upper MinAfterSunset Temperature
1 0.98 0.01 0.93 1.00 0 -2
2 0.89 0.04 0.78 0.95 0 -1
3 0.55 0.05 0.45 0.65 0 0
4 0.16 0.03 0.10 0.23 0 1
5 0.03 0.01 0.01 0.07 0 2
```

Confidence intervals are requested with confint, using either the asymptotic normal approximation or profiling.

```
> confint(fm2, type='det')
```

```

              0.025      0.975
p(Int)          -0.1946871  0.6142292
p(MinAfterSunset) -0.3985642  0.2274722
p(Temperature)   -2.4638797 -1.3233511
> confint(fm2, type='det', method = "profile")
Profiling parameter 1 of 3 ... done.
Profiling parameter 2 of 3 ... done.
Profiling parameter 3 of 3 ... done.
              0.025      0.975
p(Int)          -0.1929210  0.6208837
p(MinAfterSunset) -0.4044794  0.2244221
p(Temperature)   -2.5189984 -1.3789261

```

2.4 Model selection and model fit

Model selection and multi-model inference can be implemented after organizing models using the `fitList` function.

```

> fms <- fitList('psi(.)p(.)' = fm1, 'psi(.)p(Time+Temp)' = fm2)
> modSel(fms)

              nPars    AIC  delta  AICwt cumltvWt
psi(.)p(Time+Temp)    4 356.76   0.00 1.0e+00    1.00
psi(.)p(.)             2 461.00 104.25 2.3e-23    1.00
> predict(fms, type='det', newdata = newData)
      Predicted      SE    lower    upper
1 0.98196076 0.01266193 0.9306044 0.99549474
2 0.89123189 0.04248804 0.7763166 0.95084836
3 0.55225129 0.05102660 0.4514814 0.64890493
4 0.15658708 0.03298276 0.1021713 0.23248007
5 0.02718682 0.01326263 0.0103505 0.06948653

```

The parametric bootstrap can be used to check the adequacy of model fit. Here we use a χ^2 statistic appropriate for binary data.

```

> chisq <- function(fm) {
  umf <- getData(fm)
  y <- getY(umf)
  y[y>1] <- 1
  sr <- fm@sitesRemoved
  if(length(sr)>0)
    y <- y[-sr,,drop=FALSE]
  fv <- fitted(fm, na.rm=TRUE)
  y[is.na(fv)] <- NA
  sum((y-fv)^2/(fv*(1-fv)), na.rm=TRUE)
}
> (pb <- parboot(fm2, statistic=chisq, nsim=100))
Call: parboot(object = fm2, statistic = chisq, nsim = 100)

```

```

Parametric Bootstrap Statistics:
      t0 mean(t0 - t_B) StdDev(t0 - t_B) Pr(t_B > t0)
1 356      20.2      15.6      0.0792

```

```

t_B quantiles:
      0% 2.5% 25% 50% 75% 97.5% 100%
t*1 299 306 326 334 346 371 385

```

```

t0 = Original statistic computed from data
t_B = Vector of bootstrap samples

```

We fail to reject the null hypothesis, and conclude that the model fit is adequate.

2.5 Derived parameters and empirical Bayes methods

The `parboot` function can also be used to compute confidence intervals for estimates of derived parameters, such as the proportion of sites occupied $PAO = \sum_i z_i$ where z_i is the true occurrence state at site i , which is unknown at sites where no individuals were detected. The “colext” vignette shows examples of using `parboot` to obtain confidence intervals for such derived quantities. An alternative way achieving this goal is to use empirical Bayes methods, which were introduced in `unmarked` version 0.9-5. These methods estimate the posterior distribution of the latent variable given the data and the estimates of the fixed effects (the MLEs). The mean or the mode of the estimated posterior distribution is referred to as the empirical best unbiased predictor (EBUP), which in `unmarked` can be obtained by applying the `bup` function to the estimates of the posterior distributions returned by the `ranef` function. The following code returns the estimate of PAO and a 90% confidence interval.

```
> re <- ranef(fm2)
> EBUP <- bup(re, stat="mode")
> CI <- confint(re, level=0.9)
> rbind(PAO = c(Estimate = sum(EBUP), colSums(CI)) / 130)
      Estimate      5%      95%
PAO 0.8076923 0.7384615 0.9923077
```

Note that this is similar, but slightly lower than the population-level estimate of ψ obtained above.

A plot method also exists for objects returned by `ranef`, but distributions of binary variables are not so pretty. Try it out on a fitted abundance model instead.

References

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- [4] Darryl I. MacKenzie, James D. Nichols, G. B. Lachman, S. Droege, J. A. Royle, and C. A. Langtimm. Estimating site occupancy rates when detection probabilities are less than one. *Ecology*, 83(8):2248–2255, 2002.
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